# STAT 3008 Applied Regression Analysis Tutorial 7.

# XU Yongze

# OCT 21&22&24

# Contents

1	Review Questions	2
	1.1 Confidence ellipse (Ex5 Q3)	2
	1.2 Find SSE for different models (Ex6 Q2)	4
2	Understanding parameter estimates	6
3	Aliased and Multicollinearity	7
4	Overfitted Model	8
5	Underfitted Model (having lurking variable)	9
	5.1 Property of estimates	9
	5.2 Relationship b/w mean functions of true and mis-specified model	9
6	Drawing conclusions	10
7	More on R square	10

# **1** Review Questions

#### 1.1 Confidence ellipse (Ex5 Q3)

Confidence interval - confidence region for a single parameter; Confidence ellipse - joint confidence region for two parameters.

**Question:** 

Let 
$$n = 100, \ \hat{\beta} = (\hat{\beta}_1, \hat{\beta}_2)' = (7, 1)', \ \hat{\sigma} = 2,$$
  
$$X'X = \begin{pmatrix} 3 & 2\\ 2 & 5 \end{pmatrix}.$$

Find an inequality representing the 95% confidence ellipse for  $\beta$ . Stretch a graph for the ellipse. (Hints: Try to find some points that lie on the ellipse.)

#### Answer:

 $(1 - \alpha)$  confidence ellipse:

$$\frac{(\hat{\beta}-\beta)'(X'X)(\hat{\beta}-\beta)}{(p+1)\hat{\sigma}^2} \le F(\alpha, p+1, n-p-1)$$

Here  $p + 1 = 2, \alpha = 0.05$ , so the 95% confidence region of  $\beta$  is:

$$\frac{(\beta - \hat{\beta})^{T} (X^{T} X)(\beta - \hat{\beta})}{2\sigma^{2}} \leq F(0.05, 2, 100 - 2)$$

$$\frac{\left(\binom{\beta_{1}}{\beta_{2}} - \binom{7}{1}\right)^{T} \binom{3}{2} 2}{2 \cdot 5} \binom{\beta_{1}}{\beta_{2}} - \binom{7}{1}}{2} \leq 3.0892$$

$$(\beta_{1} - 7 - \beta_{2} - 1)\binom{3}{2} 2}{5} \binom{\beta_{1} - 7}{\beta_{2} - 1} \leq 3.0892 \times 8$$

$$(3\beta_{1} + 2\beta_{2} - 23 - 2\beta_{1} + 5\beta_{2} - 19)\binom{\beta_{1} - 7}{\beta_{2} - 1} \leq 3.0892 \times 8$$

$$3\beta_{1}^{2} + 5\beta_{2}^{2} + 4\beta_{1}\beta_{2} - 46\beta_{1} - 38\beta_{2} + 180 \leq 3.0892 \times 8$$

$$3\beta_{1}^{2} + 5\beta_{2}^{2} + 4\beta_{1}\beta_{2} - 46\beta_{1} - 38\beta_{2} + 155.2864 \leq 0$$

Trick for calculation:  $\begin{pmatrix} \beta_1 - 7 & \beta_2 - 1 \end{pmatrix} \begin{pmatrix} 3 & 2 \\ 2 & 5 \end{pmatrix} \begin{pmatrix} \beta_1 - 7 \\ \beta_2 - 1 \end{pmatrix}$ 

$$\Rightarrow 3(\beta_1 - 7)^2 + 4(\beta_1 - 7)(\beta_2 - 1) + 5(\beta_2 - 1)^2 \Rightarrow 3\beta_1^2 + 5\beta_2^2 + 4\beta_1\beta_2 - (14 \times 3 + 4)\beta_1 - (4 \times 7 + 5 \times 2)\beta_2 + 3 \times 49 + 4 \times 7 + 5 \Rightarrow 3\beta_1^2 + 5\beta_2^2 + 4\beta_1\beta_2 - 46\beta_1 - 38\beta_2 + 180$$



How to find some special points and stretch a graph for the ellipse:

• Find points A1 and A2: solve

$$3(\beta_1 - 7)^2 + 4(\beta_1 - 7)(\beta_2 - 1) + 5(\beta_2 - 1)^2 - 3.0892 \times 8 = 0$$

when  $\beta_1 = 7$ . Equivalently, to solve  $5(\beta_2 - 1)^2 - 24.7136 = 0$ . So  $\beta_2 = \pm \sqrt{\frac{25.7136}{5}} + 1$ . Hence A1 = (7, 3.223223), A2 = (7, -1.223223).

• Find points B1 and B2: to find the  $\beta_2$ 's so that the equation

$$3\beta_1^2 + 5\beta_2^2 + 4\beta_1\beta_2 - 46\beta_1 - 38\beta_2 + 155.2864 = 0$$

for  $\beta_1$  has only one solution. The previous equation is equivalent to

$$3\beta_1^2 + (4\beta_2 - 46)\beta_1 + (5\beta_2^2 - 38\beta_2 + 155.2864) = 0$$

When  $\triangle = (4\beta_2^2 - 46)^2 - 12(5\beta_2^2 - 38\beta_2 + 155.2864) = -44\beta_2^2 + 88\beta_2 + 252.5632 = 0$ , i.e., $\beta_2 = \frac{-88 \pm \sqrt{88^2 + 4 \times 44 \times 252.5632}}{-88} = 3.596165$  or -1.596165, the quadratic equation has only one solution  $\beta_1 = -\frac{4\beta_2 - 46}{6} = 5.269223$  or 8.730777, respectively. Hence B1 = (5.269223, 3.596165), B2 = (8.730777, -1.5961665). • Similarly, find pints C1 and C2.

#### **1.2** Find SSE for different models (Ex6 Q2)

Let Y = (29, 34, 19, 41, 36, 36, 24, 10)', X1 = (1, 1, 3, 3, 5, 5, 7, 7)', X2 = (10, 7, 6, 3, 2, -1, -2, -5)', X3 = (2, 6, 0, 9, 7, 9, 4, 2)'.

- i Compare the regression model  $Y = \beta_0 + \beta_1 X 1 + \beta_2 X 2 + \beta_3 X 3 + e$ and  $Y = \beta_0 + \beta_2 X 2 + e$  using a F-test. Which model will you use?
- ii Test whether  $\beta_1 = \beta_2$ , i.e. both variables produce the same effect. (Hint: rewrite the model as

$$Y = \beta_0 + \beta_1 (X1 + X2) + \beta_2 X2 + \beta_3 X3 + e \tag{1}$$

and do a t-test.)

ii Do a F-test to compare between model (1) and

$$Y = \beta_0 + \beta_1 (X1 + X2) + \beta_3 X3 + e \tag{2}$$

Show numerically the F-statistic is equal to the square of the tstatistic.

Ans:

(i)  $H_0: E(Y | X) = \beta_0 + \beta_2 X_2$   $H_1: E(Y | X) = \beta_0 + \beta_2 X_2 E(Y | X) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3$ Under  $H_1$ ,  $X = \begin{pmatrix} 1 & X_1 & X_2 & X_3 \end{pmatrix} = \begin{pmatrix} 1 & 110 & 2 \\ 1 & 1 & 7 & 6 \\ 1 & 3 & 6 & 0 \\ 1 & 3 & 3 & 9 \\ 1 & 5 & 2 & 7 \\ 1 & 5 - 1 & 9 \\ 1 & 7 - 2 & 4 \\ 1 & 7 - 5 & 2 \end{pmatrix}$ 

 $SSE_{H_1} = Y^T (I - H)Y = Y^T (I - X(X^T X)^{-1} X^T)Y = 2.2742$ 

 $df_{H_1} = n - (p+1) = 8 - 4 = 4$ 

Under 
$$H_0$$
,  $X = \begin{pmatrix} 1 & X_2 \end{pmatrix} = \begin{pmatrix} 1 & 10 \\ 1 & 7 \\ 1 & 6 \\ 1 & 3 \\ 1 & 2 \\ 1 & -1 \\ 1 & -2 \\ 1 & -5 \end{pmatrix}$   
 $SSE_{H_0} = Y^T (I - H)Y = Y^T (I - X (X^T X)^{-1} X^T)Y = 656.1994$   
 $df_{H_0} = n - (p+1) = 8 - 2 = 6$   
 $F_{(0.05, 2, 4)} = 6.9443$ 

$$Obs.T.S. = \frac{\frac{SSE_{H_0} - SSE_{H_1}}{df_{H_0} - df_{H_1}}}{\frac{SSE_{H_1}}{df_{H_1}}} = 575.0883 > 6.9443$$

# Therefore, $H_0$ is rejected at 0.05 significance level.

(ii) According to equation (1),  $H_0: \beta_1 = \beta_2 \Leftrightarrow H_0: \beta_2 = 0$ . To do t-test, first look at the 'full model'(alternative model) to find out  $\hat{\beta}_i, \hat{\sigma}$  and  $se(\hat{\beta}_i)$ .

$$E(Y | X) = \beta_0 + \beta_1 (X_1 + X_2) + \beta_2 X_2 + \beta_3 X_3$$

$$X = \begin{pmatrix} 1 & X_1 + X_2 X_2 X_3 \end{pmatrix} = \begin{pmatrix} 1 & 1110 & 2 \\ 1 & 18 & 7 & 6 \\ 1 & 9 & 6 & 0 \\ 1 & 6 & 3 & 9 \\ 1 & 7 & 2 & 7 \\ 1 & 4 & -19 \\ 1 & 5 & -24 \\ 1 & 2 & -52 \end{pmatrix}$$

$$SSE = Y^T (I - H)Y = Y^T (I - X(X^T X)^{-1} X^T)Y = 2.2742$$

$$\widehat{\sigma^2} = \frac{SSE}{df} = \frac{2.2742}{8 - 4} = 0.5685$$

$$\widehat{Var}(\widehat{\beta} \mid X) = \widehat{\sigma^2}(X^T X)^{-1} = \begin{pmatrix} 6.2637 & -1.0494 & 0.5512 & -0.1537 \\ -1.0494 & 0.1842 & -0.0990 & 0.0204 \\ 0.5512 & -0.0990 & 0.0565 & -0.0101 \\ -0.1537 & 0.0204 & -0.0101 & 0.0095 \end{pmatrix}$$

$$se(\hat{\beta}_{2}) = \sqrt{Var(\hat{\beta}_{2} \mid X)} = \sqrt{0.0565} = 0.2376$$
$$\hat{\beta} = (X^{T}X)^{-1}X^{T}Y = \begin{pmatrix} -3.1673\\ 2.6774\\ -0.3620\\ 3.1373 \end{pmatrix}$$

 $t_{(0.025,4)} = 2.7764$ 

$$obs.T.S. = \frac{\hat{\beta}_2}{se(\hat{\beta}_2)} = \frac{-0.3620}{0.2376} = -1.5233$$

Therefore,  $H_0$  is not rejected at 5% significance level.

(iii)Refer to 'Ex6 Solutions'.

# 2 Understanding parameter estimates

$$\begin{split} E(Y|X) &= \beta_0 X_0 + \beta_1 X_1 + \dots + \beta_p X_p \\ E(Fuel|X) &= 154.19 - 4.23 \, Tax + 0.47 \, Dlic - 6.14 \, Income + 18.54 \log(Miles) \\ \hline \\ \text{Unit of } \beta \text{s:} \\ \text{unit of y/unit of x. (e.g. gallon/$1000 for -6.14)} \\ \hline \\ \text{Unit of } \sigma^2 \text{:} \\ \text{unit of } y|^2 \\ \hline \\ \text{Meaning of } \beta_i \text{:} \\ \text{e.g.} \\ \text{e.g.} \\ \text{Evel decreased by 6.14 gallon when Income increased by $1000 \\ \text{Evel increased by 18.54 gallon when Miles is doubled (log_22x=1+log(x))} \\ \hline \\ \text{Meaning of } \sigma^2 \text{:} \\ \text{Variability that can't be explained by the regression line.} \end{split}$$

# **3** Aliased and Multicollinearity

#### 3.1 Aliased:

Some predictors are linear combinations of other predictors  $\Rightarrow det(X'X) = 0 \Rightarrow X'X$  noninvertible

#### 3.2 Multicollinearity:

Some predictors are highly correlated.  $\Rightarrow det(X'X) \approx 0 \Rightarrow (X'X)^{-1} = \frac{adj(X'X)}{det(X'X)} \text{ unstable}$ 

Perfect multicollinearity (correlation = 1) is equivalent to Aliased

#### & Consequence:

 $>\hat{\beta}$  is unstable

>Var $(\hat{\beta})$  can be very large

Explain intuitively:

- 1. Soma=1.59-0.116 WT2+0.056 WT9+0.048 WT18
- 2. Soma=1.59-(0.116+b) WT2+(0.056+b) WT9+0.048 WT18 b DW9
- Any b in model 2 is essentially model 1!

So the estimates of coefficients are unstable and their variance can be very large.

# 4 Overfitted Model



#### Why this answer? In general,

(Over-fitted model)

True Model:  $\mathbf{Y} = \mathbf{X}_1 \mathbf{b}_1 + \boldsymbol{\varepsilon}, \quad E(\boldsymbol{\varepsilon}) = \mathbf{0}, \quad \operatorname{Var}(\boldsymbol{\varepsilon}) = \sigma^2 \mathbf{I}$ 

 $\text{Mis-specified Model:} \quad \mathbf{Y} = \mathbf{X}\mathbf{b} + \boldsymbol{\varepsilon} \quad \text{where} \quad \mathbf{X} = (\mathbf{X}_1 \ \mathbf{X}_2), \quad \mathbf{b} = \left( \begin{array}{c} \mathbf{b}_1 \\ \mathbf{b}_2 \end{array} \right)$ 

(a) What is  $E(\widehat{\mathbf{b}})$ ?

(A) 
$$E(\hat{\ell}) = E((x'x)^{-1}x'y) = (x'x)^{-1}x'E(y)$$
  
 $= (x'x)^{-1}x'(x, \beta_1) = (x'x)^{-1}x'x(\beta_1)$   
 $= {\beta_1 \choose 0} \implies E(\hat{\ell}_1) = \beta_1 \text{ and } E(\hat{\ell}_2) = 0$ 

# **5** Underfitted Model (having lurking variable)

(Underfitted Model)

True Model:  $\mathbf{Y} = \mathbf{X}\mathbf{b} + \mathbf{Z}\boldsymbol{\gamma} + \boldsymbol{\varepsilon}, \qquad E(\boldsymbol{\varepsilon}) = \mathbf{0}, \quad |Var(\boldsymbol{\varepsilon}) = \sigma^2 \mathbf{I}$ Mis-specified Model:  $\mathbf{Y} = \mathbf{X}\mathbf{b} + \boldsymbol{\varepsilon}, \qquad E(\boldsymbol{\varepsilon}) = \mathbf{0}, \quad Var(\boldsymbol{\varepsilon}) = \sigma^2 \mathbf{I}$ Let the least squares estimate of  $\mathbf{b}$  be  $\hat{\mathbf{b}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}$ 

#### 5.1 **Property of estimates**

(a) Prove that  $\hat{\mathbf{b}}$  is a biased estimator of  $\mathbf{b}$  and find the bias.

(a) 
$$E(\hat{p}) = E[(x'x)^{-1}x'\hat{\gamma}] = (x'x)^{-1}x'E(\hat{\gamma})$$
  
 $= (x'x)^{-1}x'(x\hat{p} + z\hat{\chi})$   
 $= \hat{p} + c\hat{\chi}$   
When  $C = (x'x)^{-1}x'z$   
Thus,  $\hat{p}$  is a biased estimate of  $\hat{p}$   
with bias  $C\hat{\chi}$ .

#### 5.2 Relationship b/w mean functions of true and mis-specified model

True model:

$$E(Y|X = x, L = l) = \beta_0 + \beta_1 x + \delta l$$

♣ Wrong model:

$$E(Y|X = x) = E[E(Y|X = x, L = l)|X = x]$$
$$= E(\beta_0 + \beta_1 x + \delta l|X = x)$$
$$= \beta_0 + \beta_1 x + \delta E(L|X = x)$$

If  $E(L|X = x) = \gamma_0 + \gamma_1 x$ ,  $E(Y|X = x) = (\beta_0 + \delta\gamma_0) + (\beta_1 + \delta\gamma_0)x$ .

#### Example (Ex7 Q3)

Suppose we fit a regression with the true mean function

$$E(Y|X_1 = x_1, X_2 = x_2) = 3 + 4x_1 + 2x_2$$

Provide conditions under which the mean function for  $E(Y|X_1 = x_1)$  is linear but has a negative coefficient for  $x_1$ .

Ans:

$$\begin{split} E(Y \mid X_1 = x_1, X_2 = x_2) &= 3 + 4x_1 + 2x_2 \\ E(Y \mid X_1 = x_1) &= 3 + 4x_1 + 2E(X_2 \mid X_1 = x_1) \\ \text{Therefore, the above expectation is linear if } E(X_2 \mid X_1 = x_1) &= a + bx_1 \\ E(Y \mid X_1 = x_1) &= 3 + 4x_1 + 2(a + bx_1) \\ &= 3 + 2a + (4 + 2b)x_1 \\ \text{And the coefficient of } x_1 \text{ is negative if } 4 + 2b < 0 \rightarrow b < -2 . \end{split}$$

Conclusion: estimated parameter may not tell the true effect of a variable.

#### 6 Drawing conclusions

♣ Observational ↔ experimental
 ⇔ cannot control the value of predictors ↔ can control

Drawing conclusions:
 Observational studies - association
 Experiments - causal relationship.

# 7 More on R square

- $R^2$  tends to be large if the X are dispersed, so try best to get a sample with larger range of X
- R<sup>2</sup> is useful to measure goodness of fit ⇔ scatter plot looks like a sample from a bivariate normal distribution (elliptical shapled), i.e. without leverage point, curve, lurking variable etc.