

## 4. 1.3 Berkeley Guidance Study

- Remark:
  - Aliased:
    - Some predictors are linear combinations of other predictors
  - Multicollinearity:
    - Some predictors are highly correlated.
    - Perfect multicollinearity (correlation=1) is equivalent to being aliased.
  - Mathematically,
    - When two columns of Matrix X are similar (correlated predictors),  $(X'X)^{-1}$  is unstable.
    - Reason:  $(X'X)$ 's determinant is close to 0.
    - Therefore
      - Estimator of  $\beta$  (i.e.,  $(X'X)^{-1}X'Y$ ) is unstable
      - Variance  $\sigma^2(X'X)^{-1}$  can be very large
  - Intuitively,
    1.  $Soma = 1.59 - 0.116 \text{ WT2} + 0.056 \text{ WT9} + 0.048 \text{ WT18}$
    2.  $Soma = 1.59 - (0.116 + b) \text{ WT2} + (0.056 + b) \text{ WT9} + 0.048 \text{ WT18} - b \text{ DW9}$
    - Any  $b$  in model 2 is essentially model 1!
    - Therefore
      - Estimator of  $\beta$ ,  $(X'X)^{-1}X'Y$  is unstable
      - Variance  $\sigma^2(X'X)^{-1}$  can be very large

$$X = \begin{pmatrix} x_1 & x_2 & \dots & z \\ \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \ddots & \vdots \\ \vdots & \vdots & \ddots & \vdots \end{pmatrix}$$

$$z = x_1 + x_2$$

$$\beta = (X'X)^{-1}X'Y$$

• Conclusion: Check for multicollinearity by studying predictors' correlation!

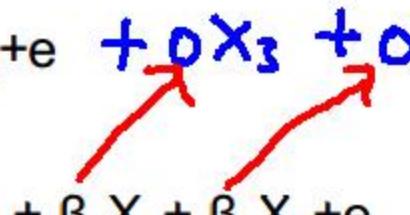
## 4.1.6 Dropping terms

- What happens when a bigger model is fit to the data from a smaller model?

- Data:

- $Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + e$  

- Model:

- $Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_3 + \beta_4 X_4 + e$  

*Answer:  $\hat{\beta}_3 \approx 0, \hat{\beta}_4 \approx 0$*

- What happens when a smaller model is fit to the data from a bigger model?

# Conditional Expectation

- Tower property

$$E(E(Y | X_1, X_2) | X_2) = E(Y | X_2)$$

Proof:  
*(optional)*  $E(E(Y | X_1, X_2) | X_2)$

$$\begin{aligned} &= E\left(\int \frac{y f_{Y, X_1, X_2}(y, X_1, X_2)}{f_{X_1, X_2}(X_1, X_2)} dy \middle| X_2\right) \\ &= \int \frac{f_{X_1, X_2}(X_1, X_2)}{f_{X_2}(X_2)} \int \frac{y f_{Y, X_1, X_2}(y, X_1, X_2)}{f_{X_1, X_2}(X_1, X_2)} dy dx_1 \\ &= \int \frac{y f_{Y, X_2}(y, X_2)}{f_{X_2}(X_2)} dy \\ &= E(Y | X_2) \end{aligned}$$

## 4.1.6 Dropping terms

- Observational analysis

- Variables are observed via sampling.
- Beyond the control of experimenter.
- Cannot avoid **lurking variable**
  - Variables that are useful but are ignored in the regression model
- True relationship:

$$E(Y | X = x, L = l) = \beta_0 + \beta_1 x + \delta l$$

- Wrong model used (L is lurking variable)

$$\begin{aligned} E(Y | X = x) &= \beta_0 + \beta_1 x + \underline{\delta E(L | X = x)} \\ &= (\beta_0 + \delta\gamma_0) + (\beta_1 + \delta\gamma_0)x \quad \text{if } E(L | X = x) = \gamma_0 + \gamma_1 x \end{aligned}$$

$$E(\underline{\beta_0 + \beta_1 x + \delta l} | X = x)$$

## Example of lurking variable

- Y: Maximum running speed (m/s)
- X: Weight (lbs)
- L: Height (cm)

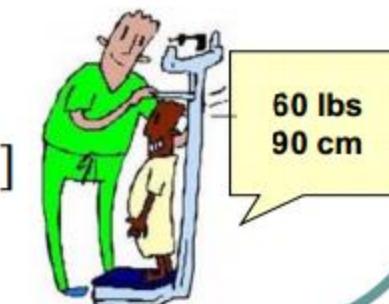
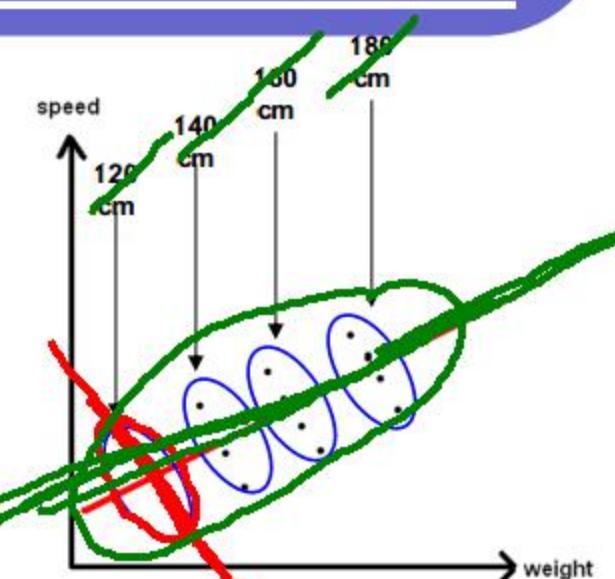
### True model

$$E(Y | X = w, L = h) = 2 - 0.05w + 0.06h$$

### Wrong model used

$$\begin{aligned} E(Y | X = w) &= 2 - 0.05w + 0.06E(L | X = w) \\ &= 2 + 0.04w \quad [\text{if } E(L | X = w) = 1.5w] \end{aligned}$$

*estimated*



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## 4.2 Experimentation v.s. Observation

- In observational study

- May have unknown effect of lurking variables.

- Can't draw causal conclusion

- e.g. Cannot say "Higher weight cause people to run faster!" X

- Only can say about association

- e.g. Can say "High weight is associated with high speed"



- In experiments

*Speed v.s. weight*

- Have control on every aspect.

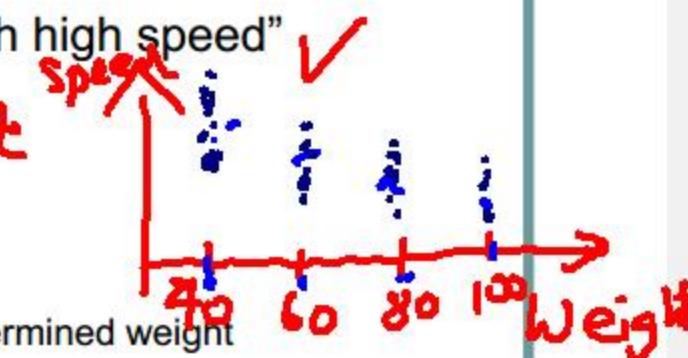
- e.g.

- Randomly assign people to achieve some pre-determined weight
- Measure their running speed.

- Lurking variables' effect are averaged out by random assignment.

- Can draw causal conclusion

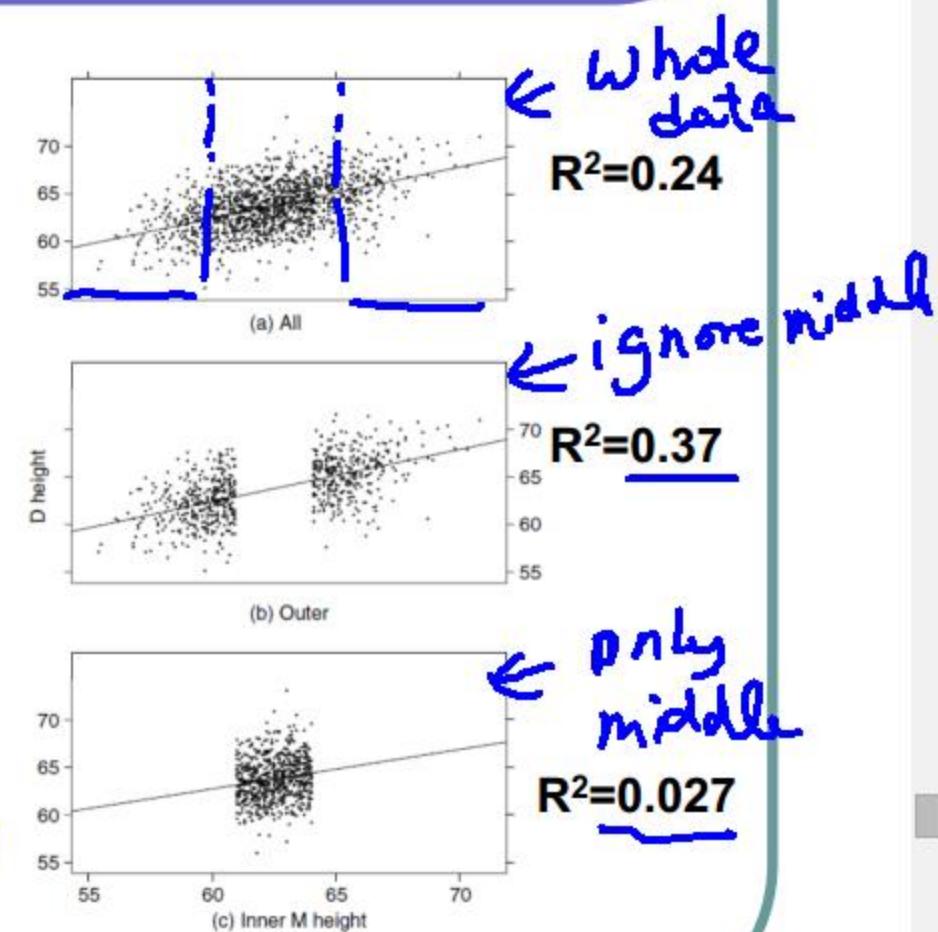
- e.g. "higher weight causes lower speed"



## 4.4 More on R<sup>2</sup>

- R<sup>2</sup> tends to be large if the X are **dispersed**
- R<sup>2</sup> tends to be small if the X are **concentrated**
- Therefore, need to be careful about sampling!

$$\begin{aligned} \sum Y \bar{Y} &= SS_{reg} + RSS \\ \sum(Y - \bar{Y})^2 &= \sum(\hat{Y} - \bar{Y})^2 + \sum(Y - \hat{Y})^2 \\ R^2 &= \frac{SS_{reg}}{\sum Y \bar{Y}} \in [0, 1] \end{aligned}$$



- $R^2$  is useful to measure goodness of regression fit if and only if the scatterplot looks like a sample from a bivariate normal distribution (elliptical shaped)

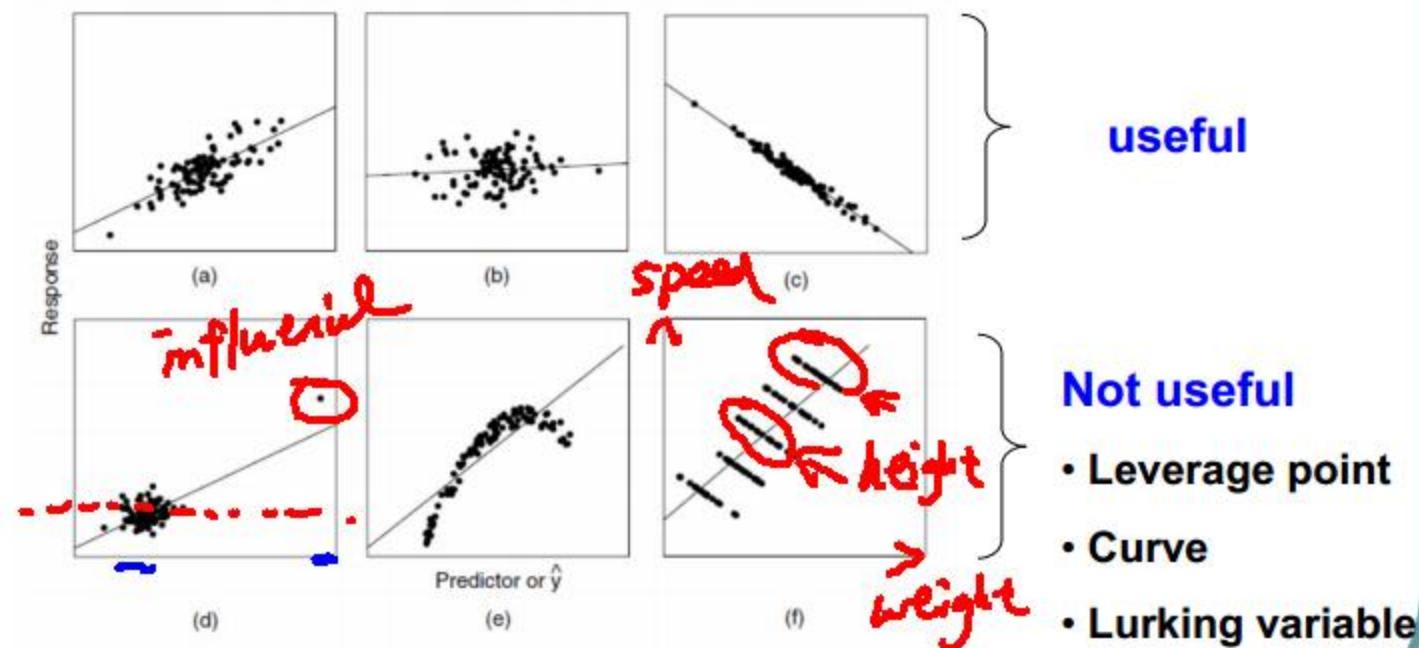


FIG. 4.3 Six summary graphs.  $R^2$  is an appropriate measure for a–c, but inappropriate for d–f.